

Bayesian nonparametric dependent model for the study of diversity for species data

Journées jeunes probabilistes et statisticiens

J. Arbel, K. Mengersen, J. Rousseau, C. King, B. Raymond
julyan.arbel@carloalberto.org

Moncalieri, Italy

April 10, 2014

Collegio Carlo Alberto



Project's bio

Authors

- Judith Rousseau (ENSAE, Université Paris-Dauphine, CREST, Paris)
- Kerrie L. Mengersen (Mathematical Sciences, Queensland University of Technology, Brisbane)
- Cath King (Australian Antarctic Division, Kingston, Tasmania 7050)
- Ben Raymond (Australian Antarctic Division, Kingston, Tasmania 7050)

Status

Chapter of my PhD thesis. Two submitted manuscripts Arbel et al. (2013b) and Arbel et al. (2013a).



Table of Contents

- 1 Ecological data and diversity
- 2 Dependent model for species data and diversity
- 3 Applications

Table of Contents

- 1 Ecological data and diversity
- 2 Dependent model for species data and diversity
- 3 Applications

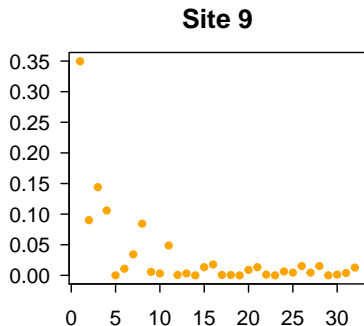
Context in ecology

- Series of measurements at different places around Casey Station, permanent base in Antarctica



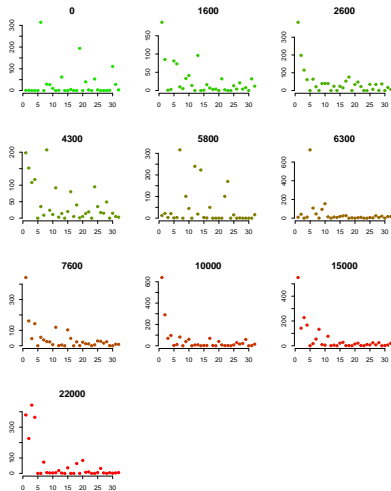
Context in ecology

- Series of measurements at different places around Casey Station, permanent base in Antarctica
- At each site: pollution level (total petroleum hydrocarbon (TPH) in mg/kg of soil), and abundance of microbes.



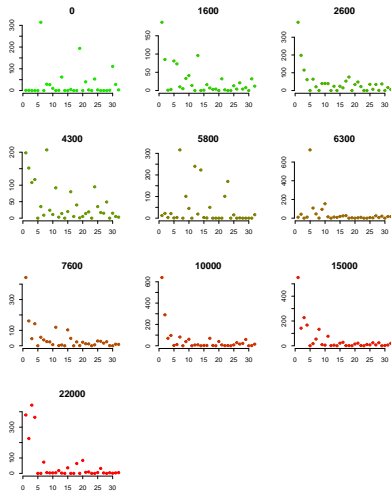
Context in ecology

- Series of measurements at different places around Casey Station, permanent base in Antarctica
- At each site: pollution level (total petroleum hydrocarbon (TPH) in mg/kg of soil), and abundance of microbes.



Context in ecology

- Series of measurements at different places around Casey Station, permanent base in Antarctica
- At each site: pollution level (total petroleum hydrocarbon (TPH) in mg/kg of soil), and abundance of microbes.
- Goals: *Assess the impact of a pollutant on the soil composition / biodiversity*, e.g. compute *effective concentration* values at level $x\%$, EC_x



Data collection

Soil samples were collected from a range of sites across a fuel contamination gradient at Australias Casey Station in East Antarctica (110° 32' E, 66° 17' S). The data comprise counts of a large number (of the order of 1 800) of microbial taxa, referred to as OTUs (operational taxonomic units; see [Schloss et al., 2009](#)), collected at 60 sites, across a range of hydrocarbon contamination ([Siciliano et al., 2014](#)). **Genomic DNA** extracted from samples was sequenced on a **454 Titanium FLX+ instrument** (Roche, Brandford, CT, USA) at the Research and Testing facility (Lubbock, TX, USA) using the **universal bacterial primers 28F and 519R** ([Dowd et al., 2008](#)). **Pyrosequencing data** were processed using the **mothur software package** ([Schloss et al., 2009](#)). This involved removal of short reads (<150bp), excessive **homopolymeric reads** (>8bp repeats) and **denoising with AmpliconNoise** (min/max flows 360/720) ([Quince et al., 2011](#)). **Preclustering** at 1% was performed to negate the per base error rate of the 454 platforms. Seed sequences were then aligned to the **SILVA 16S rRNA** gene database alignment using a **NAST alignment algorithm** ([Pruesse et al., 2007](#); [Caporaso et al., 2010](#)). Reads were then **chimaera-checked** ([Edgar et al., 2011](#)) and clustered into OTUs at 96% sequence similarity to achieve approximately species-level units as derived by [Kim et al. \(2011\)](#). Seed sequences from each OTU were then classified using a **Naïve Bayesian classifier in mothur** against the **Greengenes 16S** reference database (October 2012 version, see [McDonald et al., 2012](#)).

Diversity indices

Shannon index

Simpson index

$$H_{\text{Shan}}(\mathbf{p}) = - \sum p_j \log p_j \quad H_{\text{Simp}}(\mathbf{p}) = 1 - \sum p_j^2$$

Good index

$$H_{\text{Good},\alpha,\beta}(\mathbf{p}) = - \sum p_j^\alpha \log^\beta p_j$$

Diversity indices of microbial data

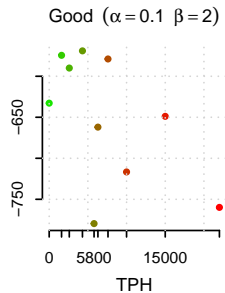
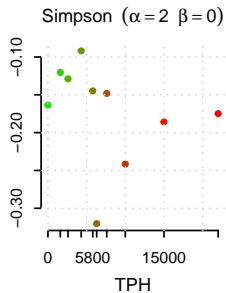
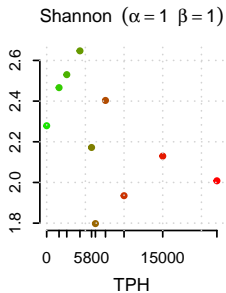


Table of Contents

- 1 Ecological data and diversity
- 2 Dependent model for species data and diversity
- 3 Applications

Model

Model the distribution of microbes in the soil:

The n th observation at site i is species j with probability $p_j(X_i)$

Model

Model the distribution of microbes in the soil:

The n th observation at site i is species j with probability $p_j(X_i)$

Model

For every site i

$$Y_{n,i} \mid \mathbf{p}(X_i), X_i \stackrel{\text{ind}}{\sim} \sum_{j=1}^{\infty} p_j(X_i) \delta_j$$

Model

Model the distribution of microbes in the soil:

The n th observation at site i is species j with probability $p_j(X_i)$

Model

For every site i

$$Y_{n,i} \mid \mathbf{p}(X_i), X_i \stackrel{\text{ind}}{\sim} \sum_{j=1}^{\infty} p_j(X_i) \delta_j$$

Parameters: $\mathbf{p} = (\mathbf{p}(X_1), \dots, \mathbf{p}(X_I)) = (p_j(X_i))_{i,j}$

Model

Model the distribution of microbes in the soil:

The n th observation at site i is species j with probability $p_j(X_i)$

Model

For every site i

$$Y_{n,i} \mid \mathbf{p}(X_i), X_i \stackrel{\text{ind}}{\sim} \sum_{j=1}^{\infty} p_j(X_i) \delta_j$$

Parameters: $\mathbf{p} = (\mathbf{p}(X_1), \dots, \mathbf{p}(X_I)) = (p_j(X_i))_{i,j}$

- Holmes et al. (2012): Dirichlet Multinomial Mixtures: Generative Models for Microbial Metagenomics (PloS one)

Model

Model the distribution of microbes in the soil:

The n th observation at site i is species j with probability $p_j(X_i)$

Model

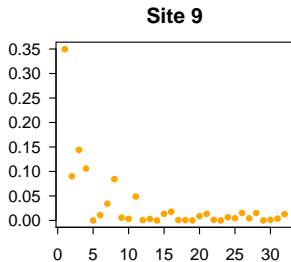
For every site i

$$Y_{n,i} \mid \mathbf{p}(X_i), X_i \stackrel{\text{ind}}{\sim} \sum_{j=1}^{\infty} p_j(X_i) \delta_j$$

Parameters: $\mathbf{p} = (\mathbf{p}(X_1), \dots, \mathbf{p}(X_I)) = (p_j(X_i))_{i,j}$

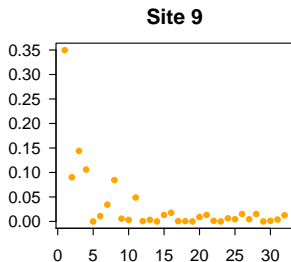
- Holmes et al. (2012): Dirichlet Multinomial Mixtures: Generative Models for Microbial Metagenomics (PloS one)
- Lijoi et al. (2007): Bayesian nonparametric estimation of the probability of discovering new species (Biometrika)

Randomizing the weights $p_j(X_i)$



Randomizing the weights $p_j(X_i)$

- Use the distribution of the weights in a Dirichlet process, obtained by a **stick-breaking** construction



Randomizing the weights $p_j(X_i)$

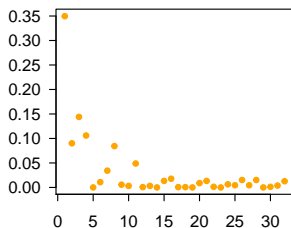
- Use the distribution of the weights in a Dirichlet process, obtained by a **stick-breaking** construction

Stick-breaking construction

$$p_1 = V_1, \quad p_j = V_j \prod_{l < j} (1 - V_l),$$

with $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$.

Site 9



Randomizing the weights $p_j(X_i)$

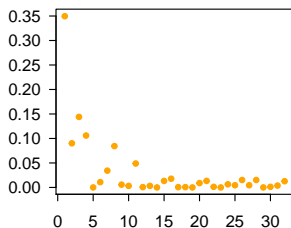
- Use the distribution of the weights in a Dirichlet process, obtained by a **stick-breaking** construction

Stick-breaking construction

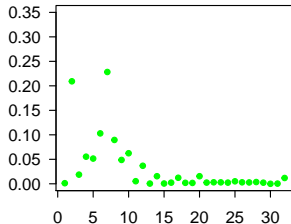
$$p_1 = V_1, \quad p_j = V_j \prod_{l < j} (1 - V_l),$$

with $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$.

Site 9



M=6



Randomizing the weights $p_j(X_i)$

- Use the distribution of the weights in a Dirichlet process, obtained by a **stick-breaking** construction

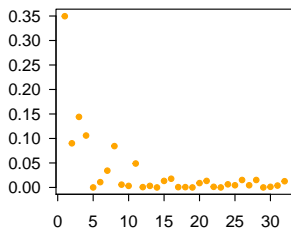
Stick-breaking construction

$$p_1 = V_1, \quad p_j = V_j \prod_{l < j} (1 - V_l),$$

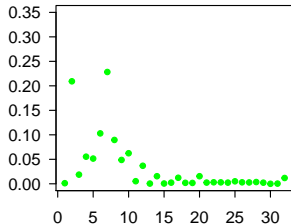
with $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$.

It is denoted $\mathbf{p} \sim \text{GEM}(M)$.

Site 9



M=6



On convergence rates

(

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

which is shown to be $\text{Ga}(N, M)$ when $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

which is shown to be $\text{Ga}(N, M)$ when

$$V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$$

Trickier when

$$V_j \stackrel{\text{ind}}{\sim} \text{Beta}(a, b + cj), a \neq 1$$

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

which is shown to be $\text{Ga}(N, M)$ when $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$

Trickier when

$$V_j \stackrel{\text{iid}}{\sim} \text{Beta}(a, b + cj), a \neq 1$$

Need a central limit theorem for

$$\sum_{j=1}^N -\log(1 - V_j)$$

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

which is shown to be $\text{Ga}(N, M)$ when $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$

Trickier when

$$V_j \stackrel{\text{iid}}{\sim} \text{Beta}(a, b + cj), a \neq 1$$

Need a central limit theorem for

$$\sum_{j=1}^N -\log(1 - V_j)$$

Limit distribution seems to be **Gumbel**...

On convergence rates

(Need control in probability tail sums

$$\sum_{j=N+1}^{\infty} p_j$$

or partial product or partial sum

$$\prod_{j=1}^N (1 - V_j)$$

$$\sum_{j=1}^N -\log(1 - V_j)$$

which is shown to be $\text{Ga}(N, M)$ when $V_j \stackrel{\text{iid}}{\sim} \text{Beta}(1, M)$

Trickier when

$$V_j \stackrel{\text{iid}}{\sim} \text{Beta}(a, b + cj), a \neq 1$$

Need a central limit theorem for

$$\sum_{j=1}^N -\log(1 - V_j)$$

Limit distribution seems to be **Gumbel...**)

Construction of the prior

- With the strick-breaking relation, a Dep – GEM prior is obtained from a Beta process.

Construction of the prior

- With the strick-breaking relation, a Dep – GEM prior is obtained from a **Beta process**.
- Such a dependent Beta process is obtained by a transformed **Gaussian process** (Rasmussen and Williams, 2006)
→ Denote by $Z \sim N(0, \sigma^2)$ a Gaussian random variable, by Φ_{σ_Z} its CDF and by F_M a $\text{Beta}(1, M)$ CDF. Then:

$$\Phi_{\sigma_Z}(Z) \sim \text{Unif}(0, 1) \text{ and } V = F_M^{-1} \circ \Phi_{\sigma_Z}(Z) \sim \text{Beta}(1, M),$$

Construction of the prior

- With the strick-breaking relation, a Dep – GEM prior is obtained from a **Beta process**.
- Such a dependent Beta process is obtained by a transformed **Gaussian process** (Rasmussen and Williams, 2006)
 → Denote by $Z \sim N(0, \sigma^2)$ a Gaussian random variable, by Φ_{σ_Z} its CDF and by F_M a Beta(1, M) CDF. Then:

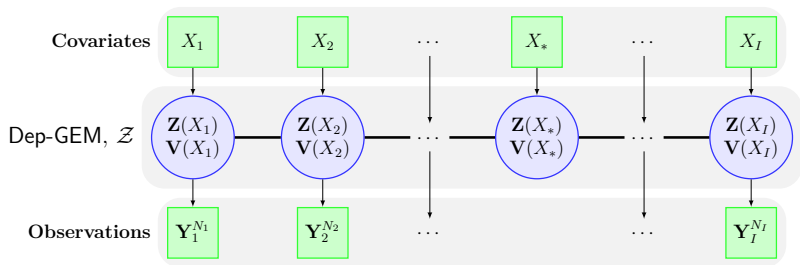
$$\Phi_{\sigma_Z}(Z) \sim \text{Unif}(0, 1) \text{ and } V = F_M^{-1} \circ \Phi_{\sigma_Z}(Z) \sim \text{Beta}(1, M),$$

- Dependence specified by covariance function

$$K(X_i, X_j) = \text{Cov}(\mathcal{Z}(X_i), \mathcal{Z}(X_j)).$$

| Covariance function | $\tilde{K}_\lambda(X_1, X_2)$ |
|--------------------------|---|
| Squared Exponential (SE) | $\exp(- (X_1 - X_2)^2 / (2\lambda^2))$ |
| Ornstein-Uhlenbeck (OU) | $\exp(- X_1 - X_2 / \lambda)$ |
| Rational Quadratic (RQ) | $(1 + (X_1 - X_2)^2 / (2\lambda^2))^{-1}$ |

Graphical model representation for the Dep – GEM model



Algorithm: Metropolis within Gibbs

Algorithm 1 Dep – GEM algorithm (Gibbs)

- 1: Update \mathbf{Z} given $(\sigma_{\mathbf{Z}}, \lambda, M)$
 - 2: Update $\sigma_{\mathbf{Z}}$ given (\mathbf{Z}, λ, M)
 - 3: Update λ given $(\mathbf{Z}, \sigma_{\mathbf{Z}}, M)$
 - 4: Update M given $(\mathbf{Z}, \sigma_{\mathbf{Z}}, \lambda)$
-

Algorithm: Metropolis within Gibbs

Algorithm 3 Dep – GEM algorithm (Gibbs)

- 1: Update \mathbf{Z} given $(\sigma_{\mathbf{Z}}, \lambda, M)$
 - 2: Update $\sigma_{\mathbf{Z}}$ given (\mathbf{Z}, λ, M)
 - 3: Update λ given $(\mathbf{Z}, \sigma_{\mathbf{Z}}, M)$
 - 4: Update M given $(\mathbf{Z}, \sigma_{\mathbf{Z}}, \lambda)$
-

Algorithm 4 MH algorithm

- 1: Given θ , propose $\theta' \sim Q_{\theta}(\cdot | \theta)$
 - 2: Compute $\rho_{\theta} = \frac{P_{\theta}(\theta') Q_{\theta}(\theta | \theta')}{P_{\theta}(\theta) Q_{\theta}(\theta' | \theta)}$
 - 3: Accept θ' w.p. $\min(\rho_{\theta}, 1)$, otherwise keep θ
-

Predictive distribution

- Predictive distribution of \mathbf{Z}_* obtained by integrating out \mathbf{Z} in the **conditional distribution** according to the **posterior distribution** $\pi(\mathbf{Z}|Y, X)$:

$$\pi(\mathbf{Z}_* | X_*, Y) = \int \pi(\mathbf{Z}_* | X_*, X, \mathbf{Z})\pi(\mathbf{Z}|Y, X)d\mathbf{Z}.$$

- No particular computational burden:

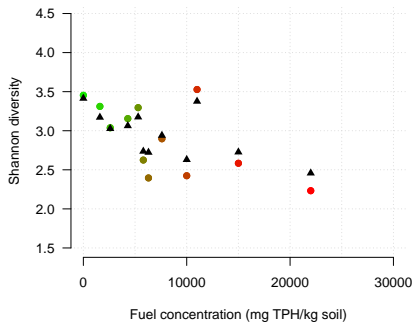
Algorithm 5 Predictive distribution simulation

- 1: Sample \mathbf{Z} from the **posterior distribution** $\pi(\mathbf{Z}|Y, X)$
 - 2: Given \mathbf{Z} , sample \mathbf{Z}_* from the **conditional distribution** $\pi(\mathbf{Z}_* | X_*, X, \mathbf{Z})$
-

Table of Contents

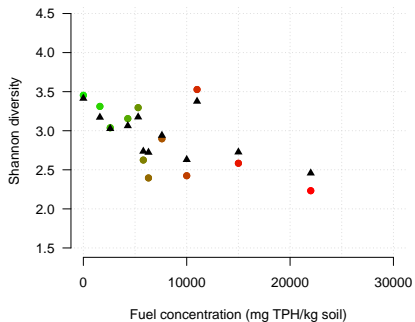
- 1 Ecological data and diversity
- 2 Dependent model for species data and diversity
- 3 Applications

Comparison of the Dep – GEM and indep. GEM models

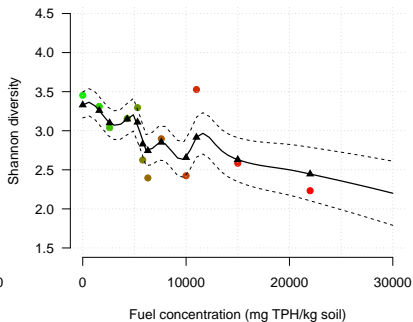


(a) GEM

Comparison of the Dep – GEM and indep. GEM models

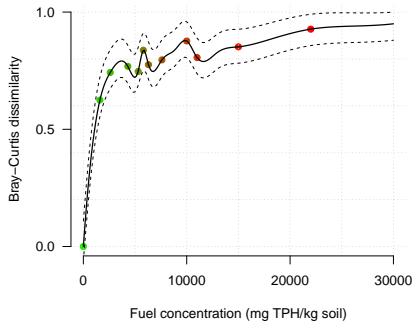


(c) GEM



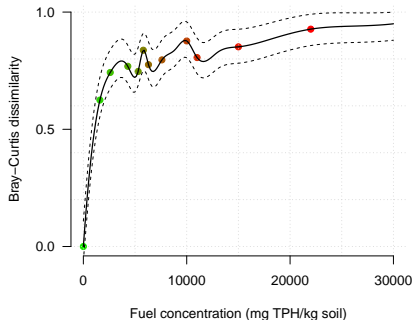
(d) Dep – GEM

Effective concentration estimation EC_x

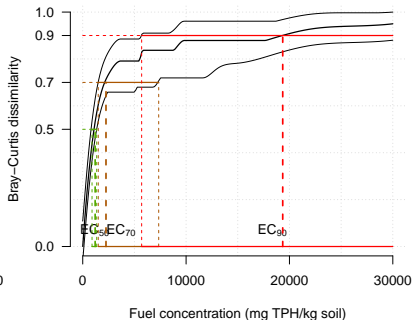


(e) Bray-Curtis dissimilarity

Effective concentration estimation EC_x



(g) Bray-Curtis dissimilarity



(h) Illustration of EC_x estimation

Future work

- Extension to multiple covariates
 - by using Gaussian random fields instead of Gaussian processes
 - model choice

Future work

- Extension to multiple covariates
 - by using Gaussian random fields instead of Gaussian processes
 - model choice
- Use of finer stick-breaking distributions
 - e.g. Beta(a, b) or Gibbs-type priors instead of Beta($1, M$)

Thank you for your attention!

- Arbel, J., Mengersen, K., Raymond, B., and King, C. (2013a). Ecotoxicological data study of diversity using a dependent Bayesian nonparametric model. *Manuscript under preparation*.
- Arbel, J., Mengersen, K., and Rousseau, J. (2013b). Bayesian nonparametric dependent models for the study of diversity in species data. *Manuscript under preparation*.
- Holmes, I., Harris, K., and Quince, C. (2012). Dirichlet Multinomial Mixtures: Generative Models for Microbial Metagenomics. *PLoS one*, 7(2):e30126.
- Lijoi, A., Mena, R. H., and Prünster, I. (2007). Bayesian nonparametric estimation of the probability of discovering new species. *Biometrika*, 94(4):769–786.
- Lijoi, A., Prünster, I., and Walker, S. G. (2008). Bayesian nonparametric estimators derived from conditional Gibbs structures. *The Annals of Applied Probability*, 18(4):1519–1547.
- Rasmussen, C. E. and Williams, C. K. I. (2006). *Gaussian Processes for Machine Learning*. MIT Press.